



Delay Analysis of Three-State Markov Channels

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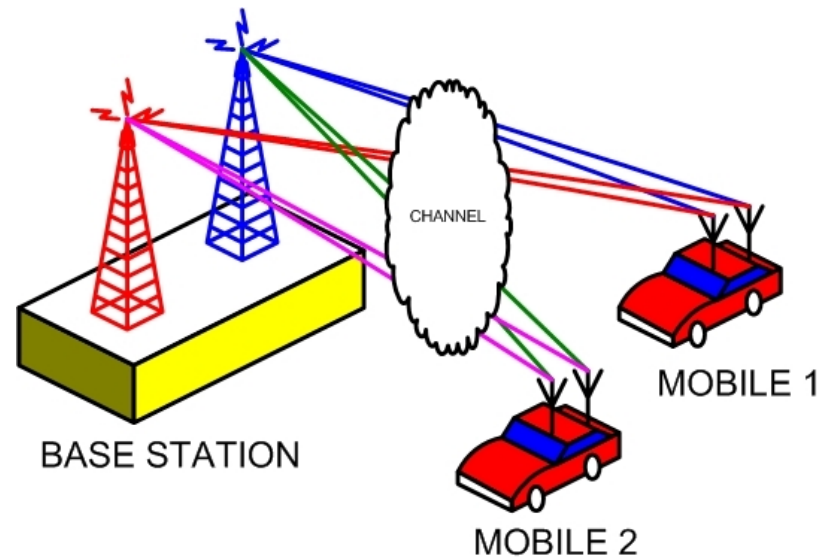


Outline

- **Background and Problem**
- Hybrid embedded Markov Chain
- Mean waiting time
- Conclusion

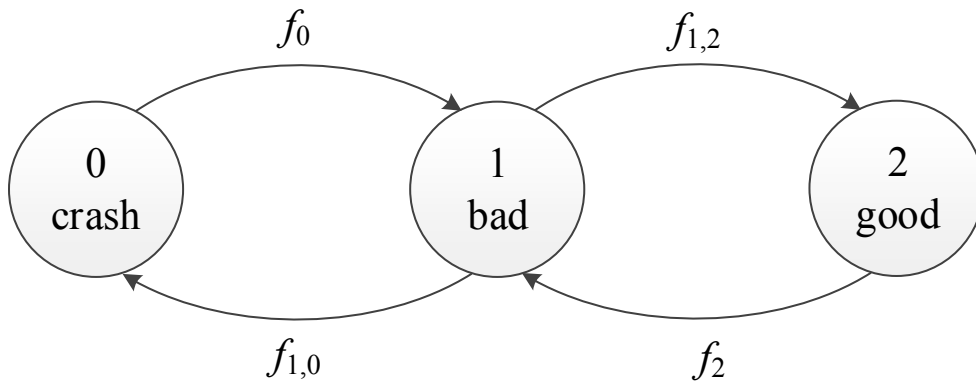
Background

- Wireless channel suffers from slow/fast fading and transmission rate changes during packet service
- Our goal is to figure out what factors affect delay performance



Three-state Markov Channel Model

- Rate fluctuation of wireless channel can be captured by a three-stage Markov Chain
 - μ_j : service rate in state j ($\mu_0 = 0, \mu_1 < \mu_2$)
 - π_j : Steady state probability that the channel is in state j

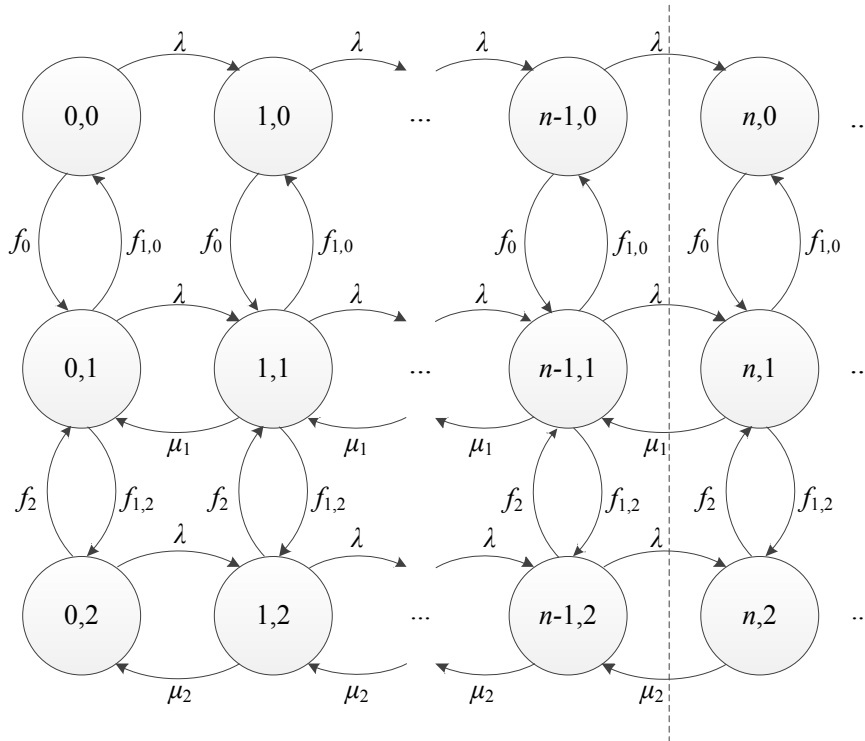


$$Q = \begin{pmatrix} -f_0 & f_0 & 0 \\ f_{1,0} & -f_1 & f_{1,2} \\ 0 & f_2 & -f_2 \end{pmatrix} \quad D = \begin{pmatrix} \mu_0 & 0 & 0 \\ 0 & \mu_1 & 0 \\ 0 & 0 & \mu_2 \end{pmatrix}$$

$$\pi Q = 0$$
$$\sum_i \pi_i = 1 \quad \Rightarrow$$
$$\pi_0 = \frac{f_{1,0}f_2}{f_0f_2 + f_2f_{1,0} + f_0f_{1,2}}$$
$$\pi_1 = \frac{f_0f_2}{f_0f_2 + f_2f_{1,0} + f_0f_{1,2}}$$
$$\pi_2 = \frac{f_0f_2}{f_0f_2 + f_2f_{1,0} + f_0f_{1,2}}$$

M/MMSP/1 queuing model

- If traffic input is a Poisson traffic with rate λ , wireless communication system can be by an M/MMSP/1 model
 - $(X(t), Y(t))$: (number of packets, channel state) at time t



Steady state probability

$$\Rightarrow p_{n,j} = \lim_{t \rightarrow \infty} P\{X(t) = n, Y(t) = j\}$$

\Rightarrow Mean Delay: An numerical Solution



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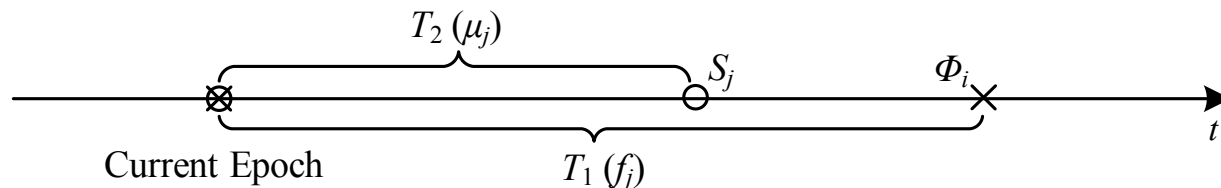
Feature of M/MMSP/1 Queue

- Service time of packets are not independently and identically distributed. It depends on the state in which the packet start being served [1].
 - Start service in larger service rate state will lead to a smaller service time.
 - The start service state of one packet is dependent on the start service state of last packet.

- We use hybrid embedded Markov chain to describe the channel state transition during the service of packets.

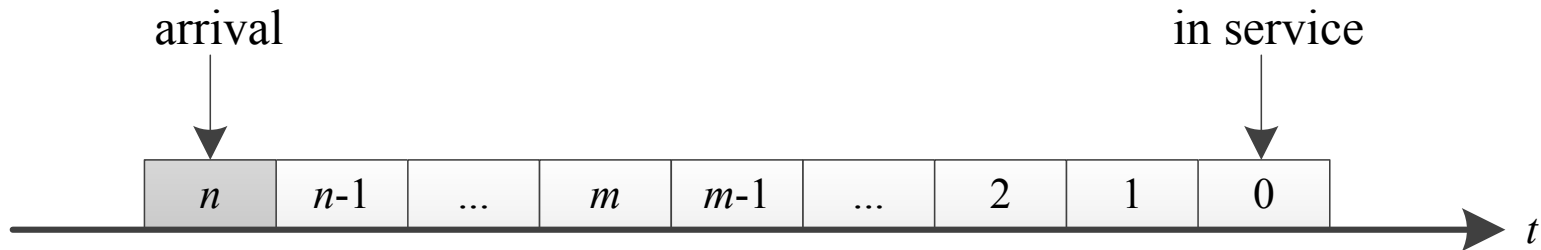
Hybrid Embedded Points

- Φ_j : Epoch when state transits, after which channel state is j
- S_j : Epoch when service starts, after which channel state is j
- If current epoch is an embedded point with state j
 - Probability that next embedded point is $\Phi_i (i \neq j)$: $f_{j,i}/(\mu_j + f_j)$;
 - Probability that next embedded point is S_j : $\mu_j/(\mu_j + f_j)$;
 - Holding time, time from current embedded point to next embedded point, follows an exponential distribution with parameter $\mu_j + f_j$.



Probability at Embedded Points

- $\hat{\pi}_{n,j}(m) = \Pr\{\text{start-service state of } m^{\text{th}} \text{ packet is } j | \text{an arrival sees } n \text{ packets in the system}\}$
- $\hat{\varphi}_{n,j}(m) = \Pr\{\text{during service of } m^{\text{th}} \text{ packet, channel transits into state } j | \text{an arrival sees } n \text{ packets in the system}\}$



Start Service Probability

- State equations of conditional start service probability

- $\hat{\pi}_{n,j}(m) = \frac{\mu_j}{\mu_j + f_j} (\hat{\pi}_{n,j}(m-1) + \hat{\varphi}_{n,j}(m-1))$

- $\hat{\varphi}_{n,j}(m) = \sum_{i \neq j} \frac{f_{i,j}}{\mu_i + f_i} (\hat{\pi}_{n,i}(m) + \hat{\varphi}_{n,i}(m))$

- Matrix form of conditional start service probability

$$\hat{\pi}_n(m) = ((D - Q)^{-1} D)^T \hat{\pi}_n(m-1)$$

where $\hat{\pi}_n(m) = (\hat{\pi}_{n,0}(m), \hat{\pi}_{n,1}(m), \hat{\pi}_{n,2}(m))^T$

- Start service probability

- $\hat{\pi}_j = \sum_{n=0}^{\infty} p_n \hat{\pi}_{n,j}(n)$

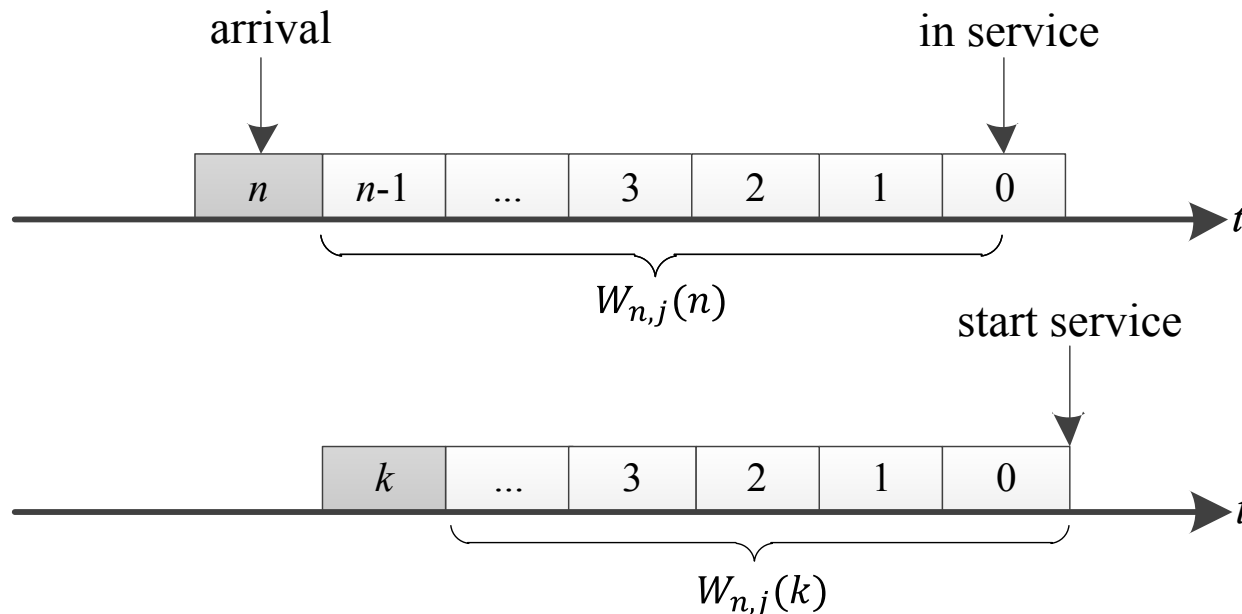


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Residual waiting time at embedded points

- A new arrival packet sees n packets in the buffer:
 - $W_{n,j}(k)$: The residual waiting time from the epoch when this new arrival becomes the k^{th} packet in the buffer while the channel is in state j to the epoch when it becomes an HOL packet.
 - $V_{n,j}(k)$: The residual waiting time from the epoch when the channel transits to state j while this new packet is the k^{th} packet in the buffer to the epoch when it becomes an HOL packet.



Mean Waiting Time

- State equations of conditional waiting time

- $$W_{n,j}(k+1) = \frac{\mu_j}{\mu_j + f_j} \left(\frac{1}{\mu_j + f_j} + W_{n,j}(k) \right) + \sum_{i \neq j} \frac{f_{j,i}}{\mu_j + f_j} \left(\frac{1}{\mu_j + f_j} + V_{n,i}(k+1) \right)$$

- $$V_{n,j}(k+1) = \frac{\mu_j}{\mu_j + f_j} \left(\frac{1}{\mu_j + f_j} + W_{n,j}(k) \right) + \sum_{i \neq j} \frac{f_{j,i}}{\mu_j + f_j} \left(\frac{1}{\mu_j + f_j} + V_{n,i}(k+1) \right)$$

- Matrix form of conditional waiting time

$$\mathbf{W}_n(k+1) = (\mathbf{D} - \mathbf{Q})^{-1} \mathbf{D} \mathbf{W}_n(k) + (\mathbf{D} - \mathbf{Q})^{-1} \mathbf{1}$$

where

$$\mathbf{W}_n(k) = \left(W_{n,0}(k), W_{n,1}(k), W_{n,2}(k) \right)^T$$

Mean Waiting Time

- Mean waiting time

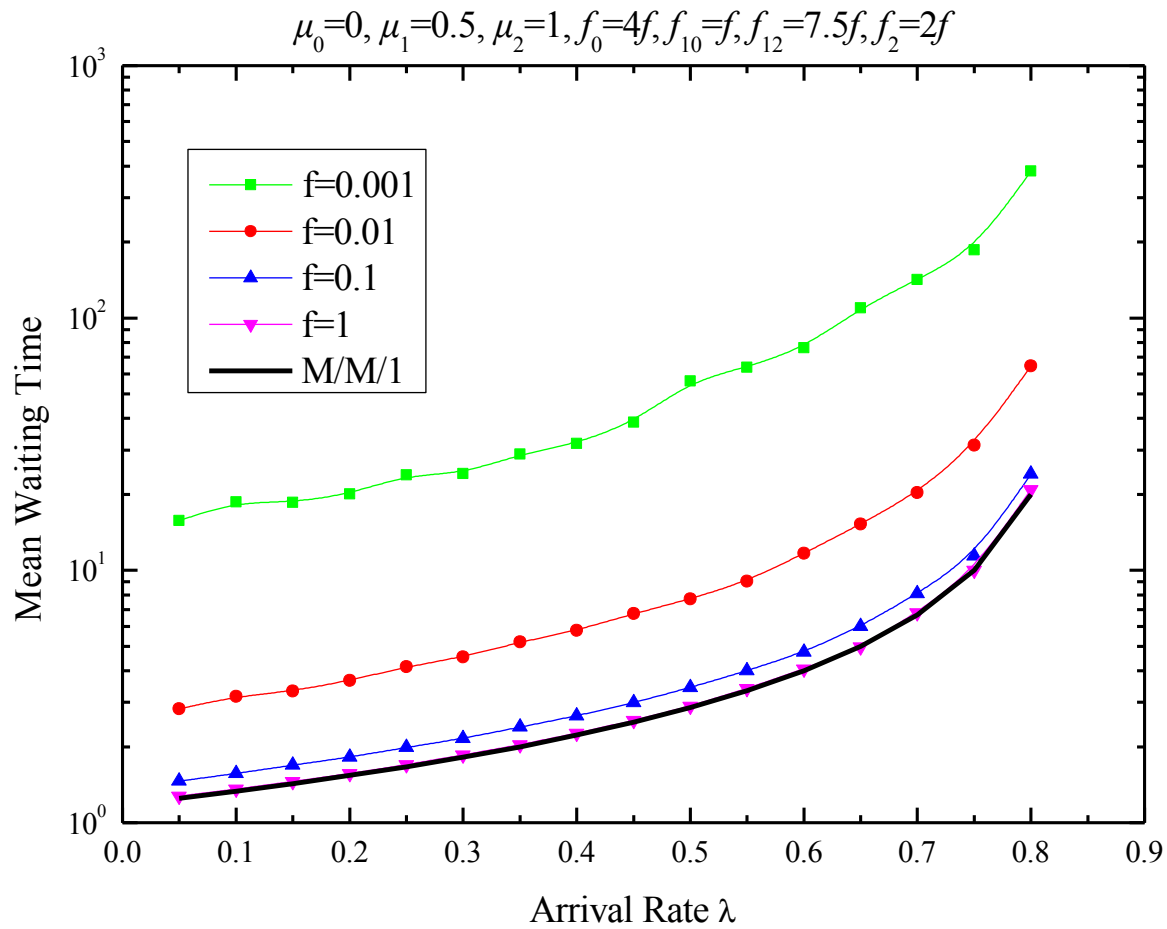
- $W = \sum_{j=0}^2 \sum_{n=1}^{\infty} W_{n,j}(n) p_{n,j}$

$$= \frac{\frac{1}{\hat{\mu}} \lambda E[T] + \frac{1}{1-\beta} \sum_{j=0}^2 E[T_j] (\pi_j - \hat{\pi}_j) - \frac{\beta}{1-\beta} \frac{1}{f_0} (\pi_0 - \hat{\pi}_0)}{1 - \frac{\lambda}{\hat{\mu}}}$$

- $\beta = \mu_1 \mu_2 / (\mu_1 \mu_2 + \mu_1 f_2 + \mu_2 f_{1,2})$ is one of the 3 eigenvalues of $(\mathbf{D} - \mathbf{Q})^{-1} \mathbf{D}$.

- $\beta \rightarrow 0$: state transition rate is much larger than service rate.
 - $\beta \rightarrow 1$: state transition rate is much smaller than service rate.

Simulation result





Conclusions

- Many problems in communication and computer networks can be modeled as M/MMSP/1 queueing model with several states, of which the numerical solution provides little physical insight
- With the help of hybrid embedded Markov chain, we obtain a structural solution and find that delay is influenced by state transition rate significantly
- Our approach can be easily extended to other finite-state M/MMSP/1 queueing model



Thank You!